

# Configurational-Bias Monte Carlo

Thijs J.H. Vlught

Professor and Chair Engineering Thermodynamics  
Delft University of Technology  
Delft, The Netherlands

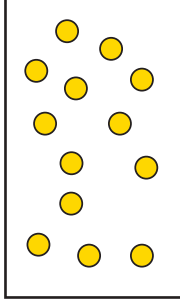
t.j.h.vlught@tudelft.nl

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## Random Sampling versus Metropolis Sampling (1)

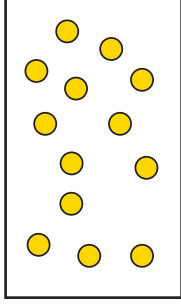
$N$  interacting particles in volume  $V$  at temperature  $T$



- vector representing positions of all particles in the system:  $\mathbf{r}^N$
- total energy:  $U(\mathbf{r}^N)$
- statistical weight of configuration  $\mathbf{r}^N$  is  $\exp[-\beta U(\mathbf{r}^N)]$  with  $\beta = 1/(k_B T)$

## Random Sampling versus Metropolis Sampling (2)

$N$  interacting particles in volume  $V$  at temperature  $T$   
pair interactions  $u(r_{ij})$



$$U(\mathbf{r}^N) = \sum_{i=1}^{N-1} \sum_{j=i+1}^N u(r_{ij}) = \sum_{i < j} u(r_{ij})$$

$$Q(N, V, T) = \frac{1}{\Lambda^{3N} N!} \int d\mathbf{r}^N \exp[-\beta U(\mathbf{r}^N)]$$

$$F(N, V, T) = -k_B T \ln Q(N, V, T)$$

## Random Sampling versus Metropolis Sampling (3)

Computing the ensemble average  $\langle \dots \rangle$  of a certain quantity  $A(\mathbf{r}^N)$

- Random Sampling of  $\mathbf{r}^N$ :

$$\langle A \rangle = \lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n A(\mathbf{r}_i^N) \exp[-\beta U(\mathbf{r}_i^N)]}{\sum_{i=1}^n \exp[-\beta U(\mathbf{r}_i^N)]}$$

Usually this leads to  $\langle A \rangle = "0" / "0" = ???$

- Metropolis sampling: generate  $n$  configurations  $\mathbf{r}^N$  with probability proportional to  $\exp[-\beta U(\mathbf{r}_i^N)]$ , therefore:

$$\langle A \rangle = \lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n A(\mathbf{r}_i^N)}{n}$$

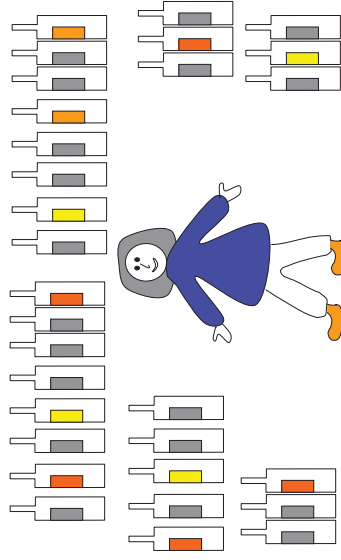
## Simulation Technique (1)

What is the ratio of red wine/white wine in the Netherlands?



## Simulation Technique (2)

Bottoms up



Simulation Technique (3)



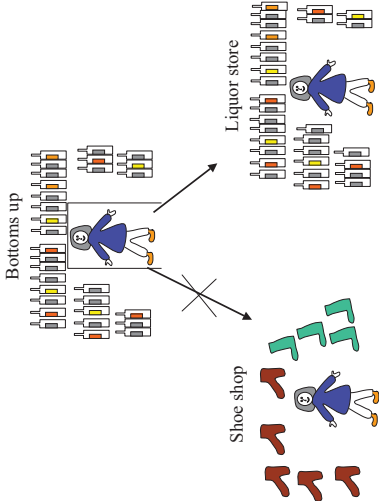
Simulation Technique (4)



Simulation Technique (5)



Simulation Technique (6)



Metropolis Monte Carlo (1)

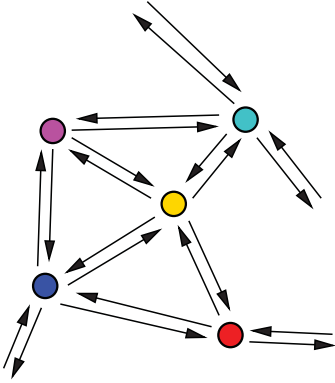
How to generate configurations  $r_i$  with a probability proportional to  $\mathcal{N}(r_i) = \exp[-\beta U(r_i)]$  ???



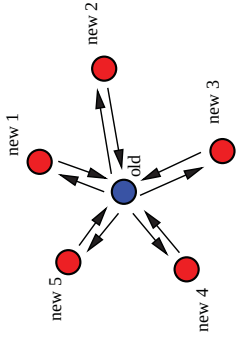
N. Metropolis, A.W. Rosenbluth, M.N. Rosenbluth, A.H. Teller and E. Teller, "Equation of State Calculations by Fast Computing Machines," J. Chem. Phys., 1953, 21, 1087-1092.

Metropolis Monte Carlo (2)

Whatever our rule is to move from one state to the next, the equilibrium distribution should not be destroyed



## Move from the old state ( $o$ ) to a new state ( $n$ ) and back



$$\text{leaving state } o = \text{entering state } o = \sum_n [\mathcal{N}(o) \alpha(o \rightarrow n) \text{acc}(o \rightarrow n)] = \sum_n [\mathcal{N}(n) \alpha(n \rightarrow o) \text{acc}(n \rightarrow o)]$$

$\mathcal{N}^{(i)}$  : probability to be in state  $i$  (here: proportional to  $\exp[-\beta U(\mathbf{r}_i)]$ )  
 $\alpha(x \rightarrow y)$  : probability to attempt move from state  $x$  to state  $y$   
 $\text{acc}(x \rightarrow y)$ : probability to accept move from state  $x$  to state  $y$

## Detailed Balance (1)

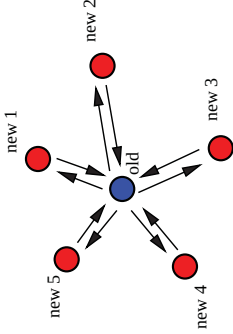
Requirement (balance):

$$\mathcal{N}(o) \sum_n [\alpha(o \rightarrow n) \text{acc}(o \rightarrow n)] = \sum_n [\mathcal{N}(n) \alpha(n \rightarrow o) \text{acc}(n \rightarrow o)]$$

**Detailed balance:** much stronger condition

$$\mathcal{N}(o) \alpha(o \rightarrow n) \text{acc}(o \rightarrow n) = \mathcal{N}(n) \alpha(n \rightarrow o) \text{acc}(n \rightarrow o)$$

for every pair  $o, n$



## Detailed Balance (2)

$$\mathcal{N}(o) \alpha(o \rightarrow n) \text{acc}(o \rightarrow n) = \mathcal{N}(n) \alpha(n \rightarrow o) \text{acc}(n \rightarrow o)$$

- $\alpha(x \rightarrow y)$ ; probability to select move from  $x$  to  $y$
- $\text{acc}(x \rightarrow y)$ ; probability to accept move from  $x$  to  $y$
- often (but not always);  $\alpha(o \rightarrow n) = \alpha(n \rightarrow o)$

Therefore (note that  $\Delta U = U(n) - U(o)$ ):

$$\frac{\text{acc}(o \rightarrow n)}{\text{acc}(n \rightarrow o)} = \frac{\alpha(n \rightarrow o) \exp[-\beta U(n)]}{\alpha(o \rightarrow n) \exp[-\beta U(o)]} = \frac{\alpha(n \rightarrow o)}{\alpha(o \rightarrow n)} \exp[-\beta \Delta U]$$

In case that  $\alpha(o \rightarrow n) = \alpha(n \rightarrow o)$

$$\frac{\text{acc}(o \rightarrow n)}{\text{acc}(n \rightarrow o)} = \exp[-\beta \Delta U]$$

## Metropolis Acceptance Rule

General:

$$\frac{\text{acc}(o \rightarrow n)}{\text{acc}(n \rightarrow o)} = X$$

Choice made by Metropolis (note: infinite number of other possibilities)

$$\text{acc}(o \rightarrow n) = \min(1, X)$$

Note that  $\min(a, b) = a$  if  $a < b$  and  $b$  otherwise

- always accept when  $X \geq 1$
- when  $X < 1$ , generate uniformly distributed random number between 0 and 1 and accept or reject according to  $\text{acc}(o \rightarrow n)$

## Monte Carlo Casino



## Smart Monte Carlo: $\alpha(o \rightarrow n) \neq \alpha(n \rightarrow o)$

Not a random displacement  $\Delta r$  uniformly from  $[-\delta, \delta]$ , but instead

$$\Delta r = r^{(\text{new})} - r^{(\text{old})} = A \times F + \delta r$$

$F$  : force on particle

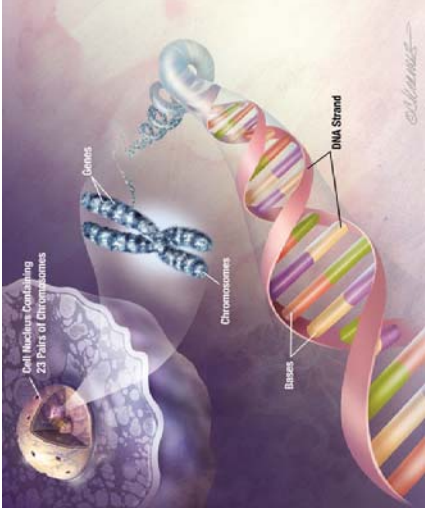
$A$  : constant

$\delta r$  : taken from Gaussian distribution with width  $2A$   
 so  $P(\delta r) \sim \exp[-(\delta r)^2/4A]$

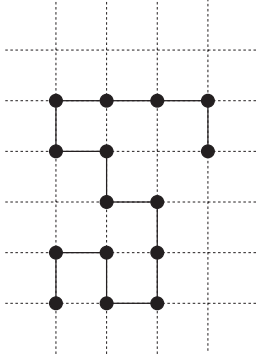
$$P(r_{\text{new}}) \sim \exp \left[ -\frac{(r_{\text{new}} - (r_{\text{old}} + A \times F(o)))^2}{4A} \right]$$

$$\frac{\alpha(o \rightarrow n)}{\alpha(n \rightarrow o)} = \frac{\exp \left[ -\frac{(\Delta r - A \times F(o))^2}{4A} \right]}{\exp \left[ -\frac{(\Delta r + A \times F(n))^2}{4A} \right]}$$

Chain Molecules



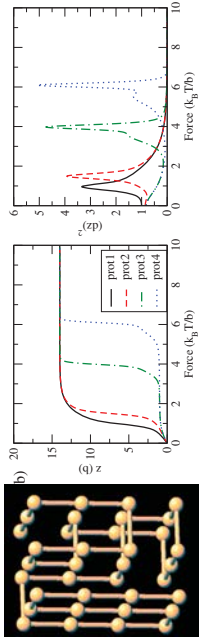
Self-Avoiding Walk on a Cubic Lattice



- 3D lattice; 6 lattice directions
- only 1 monomer per lattice site (otherwise  $U = \infty$ )
- interactions only when  $|r_{ij}| = 1$  and  $|i - j| > 1$

Simple Model for Protein Folding

20 by 20 interaction matrix  $\Delta_{ij}$   
YPDLTKWHAMEAGKIRFSVPDACLNGEGIRQVTLSN  
(E. Jarkova, T.J.H. Vlugt, N.K. Lee, J. Chem. Phys., 2005, 122, 114904)



Number of Configurations without Overlap

Random Chains:

$$\langle R \rangle = \lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n R_i \exp[-\beta U_i]}{\sum_{i=1}^n \exp[-\beta U_i]}$$

Fraction of chains without overlap decreases exponentially as a function of chainlength ( $N$ )

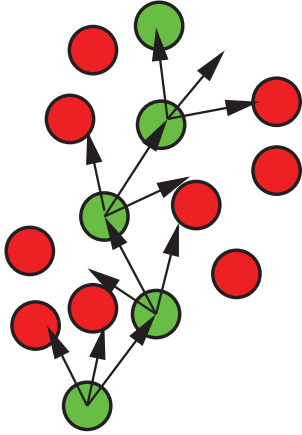
| $N$ | total ( $= 6^{N-1}$ ) | without overlap | fraction no overlap  |
|-----|-----------------------|-----------------|----------------------|
| 2   | 6                     | 6               | 1                    |
| 6   | 7776                  | 3534            | 0.454                |
| 8   | 279936                | 81390           | 0.290                |
| 10  | 10077696              | 1853886         | 0.183                |
| 12  | 362797056             | 41934150        | 0.115                |
| 13  | 2176782336            | 198842742       | 0.091                |
| 14  | 13060694016           | 943974510       | 0.072                |
| 15  | 78364164096           | 4468911678      | 0.057                |
| 16  | 470184984576          | 21175146054     | 0.045                |
| 50  | ...                   | ...             | $1.3 \times 10^{-5}$ |

Rosenbluth Sampling (1)

1. Place first monomer at a random position
2. For the next monomer ( $i$ ), generate  $k$  trial directions ( $j = 1, 2, \dots, k$ ) each with energy  $u_{ij}$
3. Select trial direction  $j^*$  with a probability 
$$P_{j^*} = \frac{\exp[-\beta u_{ij^*}]}{\sum_{j=1}^k \exp[-\beta u_{ij}]}$$
4. Continue with step 2 until the complete chain is grown ( $N$  monomers)

Rosenbluth Sampling (2)

$$P_{j^*} = \frac{\exp[-\beta u_{ij^*}]}{\sum_{j=1}^k \exp[-\beta u_{ij}]}$$



### Rosenbluth Sampling (3)

Probability to choose trial direction  $j^*$  for the  $i$  th monomer

$$P_{j^*} = \frac{\exp[-\beta u_{ij^*}]}{\sum_{j=1}^k \exp[-\beta u_{ij}]}$$

Probability to grow this chain ( $N$  monomers,  $k$  trial directions)

$$P_{\text{chain}} = \prod_{i=1}^N P_{j^*(i)} = \frac{\prod_{i=1}^N \exp[-\beta u_{ij^*(i)}]}{\prod_{i=1}^N \sum_{j=1}^k \exp[-\beta u_{ij}]} = \frac{\exp[-\beta U_{\text{chain}}]}{W_{\text{chain}}}$$

### Rosenbluth Sampling (4)

Probability to grow this chain ( $N$  monomers,  $k$  trial directions)

$$P_{\text{chain}} = \frac{\prod_{i=1}^N \exp[-\beta u_{ij^*(i)}]}{\prod_{i=1}^N \sum_{j=1}^k \exp[-\beta u_{ij}]} = \frac{\exp[-\beta U_{\text{chain}}]}{W_{\text{chain}}}$$

Therefore, weightfactor for each chain  $i$  is the Rosenbluth factor  $W_i$ :

$$\langle R \rangle_{\text{Boltzmann}} = \lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n W_i \times R_i}{\sum_{i=1}^n W_i}$$

The unweighted distribution is called the Rosenbluth distribution:

$$\langle R \rangle_{\text{Rosenbluth}} = \lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n R_i}{n}$$

Of course:  $\langle R \rangle_{\text{Rosenbluth}} \neq \langle R \rangle_{\text{Boltzmann}}$

### Intermezzo: Ensemble Averages at Different Temperatures

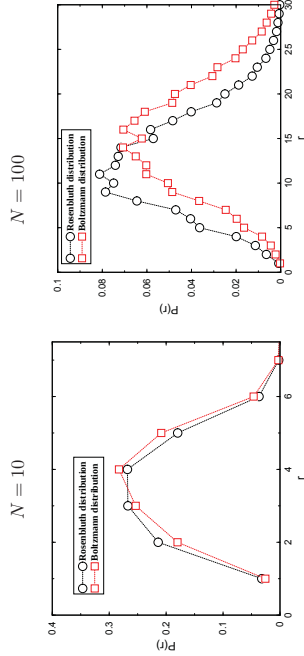
Ensemble averages at  $\beta^*$  can (in principle) be computed from simulations at  $\beta$ :

$$\begin{aligned} \langle U \rangle_{\beta} &= \frac{\int d\mathbf{r}^N U(\mathbf{r}^N) \exp[-\beta U(\mathbf{r}^N)]}{\int d\mathbf{r}^N \exp[-\beta U(\mathbf{r}^N)]} \\ &= \frac{\int d\mathbf{r}^N U(\mathbf{r}^N) \exp[-\beta^* U(\mathbf{r}^N)] \exp[(\beta^* - \beta) \times U(\mathbf{r}^N)]}{\int d\mathbf{r}^N \exp[-\beta^* U(\mathbf{r}^N)] \exp[(\beta^* - \beta) \times U(\mathbf{r}^N)]} \\ &= \frac{\langle U(\mathbf{r}^N) \exp[(\beta^* - \beta) \times U(\mathbf{r}^N)] \rangle_{\beta^*}}{\langle \exp[(\beta^* - \beta) \times U(\mathbf{r}^N)] \rangle_{\beta^*}} \\ &= \frac{\langle U(\mathbf{r}^N) \exp[\Delta\beta \times U(\mathbf{r}^N)] \rangle_{\beta^*}}{\langle \exp[\Delta\beta \times U(\mathbf{r}^N)] \rangle_{\beta^*}} \end{aligned}$$

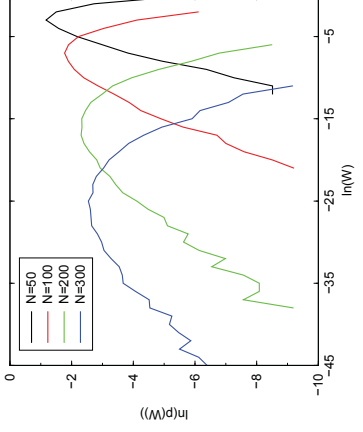
Useful or not???

### Rosenbluth Distribution Differs from Boltzmann Distribution

Probability distribution for the *end-to-end distance*  $r$



### Distribution of Rosenbluth Weights



Of course, Probability( $W = 0$ )  $\neq 0$  (not shown in this figure)

### Pruned-Enriched Rosenbluth Method (1)

Grassberger (1997): grow chains using Rosenbluth Method:

$$W = \sum_{j=1}^6 \frac{\exp[-\beta u_{2j}]}{6} \times \prod_{i=3}^5 \sum_{j=1}^5 \frac{\exp[-\beta u_{ij}]}{5} = \prod_{i=3}^5 \sum_{j=1}^5 \frac{\exp[-\beta u_{ij}]}{5}$$

Two additional elements:

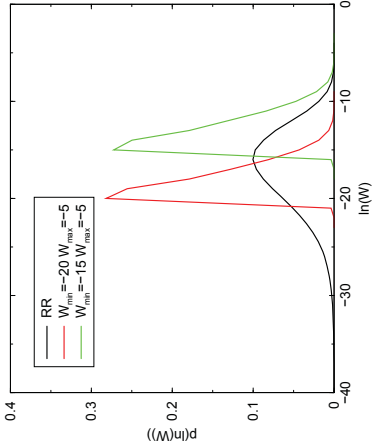
- **Enriching**  
If  $W > W_{\text{max}}$  during the construction of the chain,  $k$  copies of the chain are generated, each with a weight of  $W/k$ . This is a deterministic process. The growth of those  $k$  chains is continued.
- **Pruning**  
If  $W < W_{\text{min}}$  during the construction of a chain, with a probability of  $1/2$  the chain is pruned resulting in  $W = 0$ . If the chain survives, the Rosenbluth weight is multiplied by 2 and the growth of the chain is continued.

Pruned-Enriched Rosenbluth Method (2)



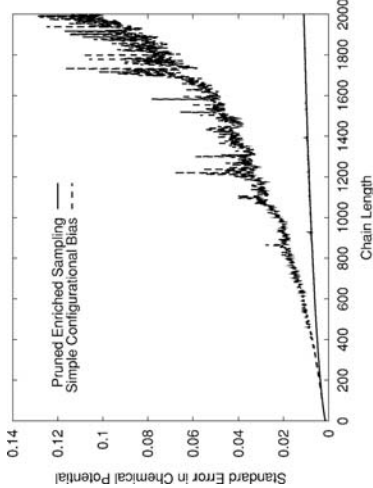
Pruned-Enriched Rosenbluth Method (3)

Example:  $N = 200$ ,  $k = 2$ ,  $\beta\mu_{\text{ex}} = -\ln\langle W \rangle = -12.14$



Pruned-Enriched Rosenbluth Method (4)

$\beta\mu_{\text{ex}} = -\ln\langle W \rangle$ , Ann. Rev. of Phys. Chem. 1999, 50, 377-411

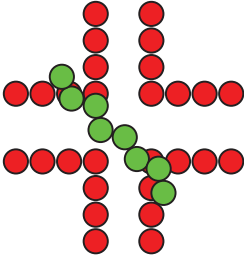


Static versus Dynamic Monte Carlo

- **Static Monte Carlo**
  - create single chain conformations, and use correct weight factor (random insertion, Rosenbluth scheme, PERM) to compute single chain averages
  - **single chain only; no Markov chain**
- **Dynamic Monte Carlo**
  - create Markov chain, accept/reject new configuration, acceptance rules should obey detailed balance
  - **multi-chain system, usable for all ensembles (e.g. Gibbs,  $\mu VT$ )**
  - Configurational-Bias Monte Carlo (CBMC); Recoil Growth (RG), Dynamic PERM (DPERM)

Random Insertion of Chains is Useless

| Chain Length | Probability without overlaps |
|--------------|------------------------------|
| 1            | $10^{-2}$                    |
| 2            | $10^{-4}$                    |
| 3            | $10^{-6}$                    |
| ...          | ...                          |
| 8            | $10^{-16}$                   |



Configurational-Bias Monte Carlo

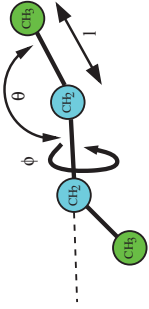
- Generate configurations using the Rosenbluth scheme
- Accept/Reject these configurations in such a way that detailed balance is obeyed
- Split potential energy into **"bonded"** (bond-stretching, bending, torsion) and **"non-bonded"** (i.e. Lennard-Jones and/or Coulombic) interactions
- Generate ( $k$ ) trial positions according to **bonded** interactions (unbranched chain:  $l, \theta, \phi$  are independent)

$$U_{\text{bonded}} = U_{\text{stretch}}(l) + U_{\text{bend}}(\theta) + U_{\text{tors}}(\phi)$$

$$\begin{aligned} P(l) &\sim dl l^2 \exp[-\beta u(l)] \\ P(\theta) &\sim d\theta \sin(\theta) \exp[-\beta u(\theta)] \\ P(\phi) &\sim d\phi \exp[-\beta u(\phi)] \end{aligned}$$



### Generate Trial Configurations: Linear Alkane



$$u(l) = (k_l/2)(l - l_0)^2$$

$$u(\theta) = (k_\theta/2)(\theta - \theta_0)^2$$

$$u(\phi) = \sum_{i=0}^5 c_i \cos^i(\phi)$$

$$P(l) \sim dl l^2 \exp[-\beta u(l)]$$

$$P(\theta) \sim d\theta \sin(\theta) \exp[-\beta u(\theta)]$$

$$P(\phi) \sim d\phi \exp[-\beta u(\phi)]$$

### Generate a Random Number from a Distribution (3)

```

subroutine bend-tors
  lready=.false
  do while (.not. lready)
    call ransphere(dx,dy,dz)
    x = xold + dx
    y = yold + dy
    z = zold + dz
    call bend(tbond, x,y,z)
    call tors(ttors, x,y,z)
    if(tranf() lt. exp(-beta*(tbond+ttors)))
      + lready=.true
  enddo
end

```

generate appropriate  $\theta$  and  $\phi$

random vector on unit sphere  
monomer position

bending energy/  
torsion energy/  
accept or reject

### Generate a Random Number from a Distribution (1)

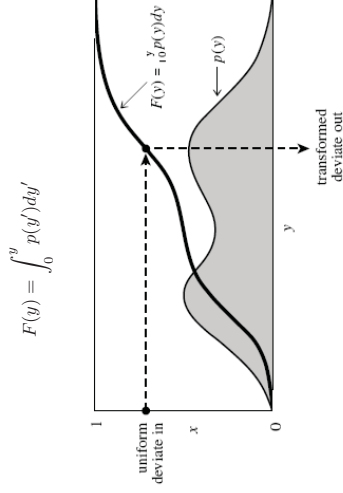


Figure 7.2.1. Transformation method for generating a random deviate  $y$  from a known probability distribution  $p(y)$ . The indefinite integral of  $p(y)$  must be known and invertible. A uniform deviate  $x$  is chosen between 0 and 1. Its corresponding  $y$  on the definite-integral curve is the desired deviate.

### Generate a Random Number from a Distribution (2)

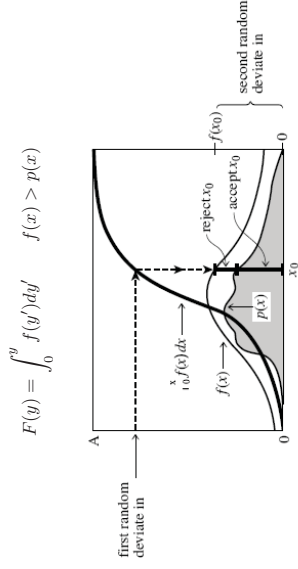
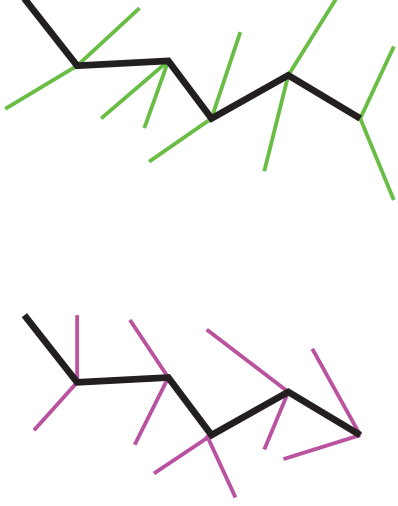


Figure 7.3.1. Rejection method for generating a random deviate  $x$  from a known probability distribution  $p(x)$  that is everywhere less than some other function  $f(x)$ . The transformation method is first used to generate a random deviate  $x$  of the distribution  $f$  (compare Figure 7.2.1). A second uniform deviate is used to decide whether to accept or reject that  $x$ . If it is rejected, a new deviate of  $f$  is found, and so on. The ratio of accepted to rejected points is the ratio of the area under  $p$  to the area between  $p$  and  $f$ .

### Super-Detailed Balance (1)

Same chain can be grow for different sets of trial directions...



### CBMC Algorithm

- Generate a trial configuration using the Rosenbluth scheme.  $k$  trial segments  $\{\mathbf{b}\}_k = \{\mathbf{b}_1 \dots \mathbf{b}_k\}$ , each trial segment is generated according to

$$P(\mathbf{b}) \sim \exp[-\beta u_{\text{bonded}}(\mathbf{b})]$$

- Compute non-bonded energy, select configuration  $i$  with a probability
 
$$P(\mathbf{b}_i) = \frac{\exp[-\beta u_{\text{non-bond}}(\mathbf{b}_i)]}{\sum_{j=1}^k \exp[-\beta u_{\text{non-bond}}(\mathbf{b}_j)]} = \frac{\exp[-\beta u_{\text{non-bond}}(\mathbf{b}_i)]}{w_i}$$

- Continue until chain is grown,  $W(n) = \prod_{l=1}^n w_l$

- Similar procedure for old configuration, generate  $k-1$  trial positions (trial position 1 is the old configuration itself), leading to  $W(o)$

- Accept/reject according to

$$\text{acc}(o \rightarrow n) = \min(1, W(n)/W(o))$$

## Super-Detailed Balance (2)

Flux of configurations

$$\begin{aligned} K(o \rightarrow n) &= \sum_{b_n, b_o} K(o \rightarrow n | b_n, b_o) \\ K(n \rightarrow o) &= \sum_{b_n, b_o} K(n \rightarrow o | b_n, b_o) \end{aligned}$$

Detailed balance requires

$$K(o \rightarrow n) = K(n \rightarrow o)$$

Possible solution (**super-detailed balance**):

$$K(o \rightarrow n | b_n, b_o) = K(n \rightarrow o | b_n, b_o)$$

for each  $b_n, b_o$ .

## Super-Detailed Balance (3)

Detailed balance for each set  $b_n, b_o$ :

$$\begin{aligned} K(o \rightarrow n | b_n, b_o) &= N(o) \times \alpha(o \rightarrow n | b_n, b_o) \times \text{acc}(o \rightarrow n | b_n, b_o) \\ &= \frac{\exp[-\beta U(o)] \times C \exp[-\beta u_{\text{bonded}}(n)] \times W(n)}{\exp[-\beta u_{\text{non-b}}(n)] \times P(b_n, b_o) \times \text{acc}(o \rightarrow n | b_n, b_o)} \\ K(n \rightarrow o | b_n, b_o) &= \frac{\exp[-\beta U(n)] \times C \exp[-\beta u_{\text{bonded}}(o)] \times W(o)}{\exp[-\beta u_{\text{non-b}}(o)] \times P(b_n, b_o) \times \text{acc}(n \rightarrow o | b_n, b_o)} \end{aligned}$$

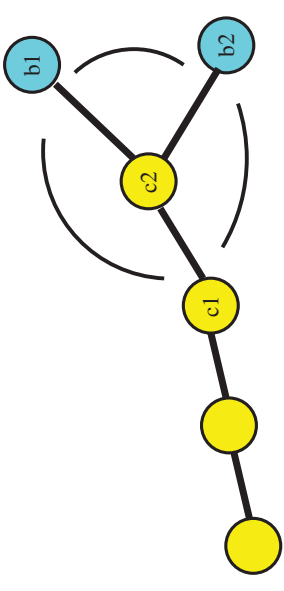
As

$$U = u_{\text{non-b}} + u_{\text{bonded}}$$

therefore

$$\frac{\text{acc}(o \rightarrow n)}{\text{acc}(n \rightarrow o)} = \frac{W(n)}{W(o)}$$

## Branched Molecules (1)



$$P(b_1, b_2) \sim \exp[-\beta[u_{\text{bond}}(c_1, c_2, b_1)] + u_{\text{bond}}(c_1, c_2, b_2)] + u_{\text{bond}}(b_1, c_2, b_2)]$$

## Branched Molecules (2)

Use CBMC to generate internal configurations:

- Generate  $n_i$  *random* trial positions and select one ( $i$ ) with a probability

$$p^{\text{int}}(i) = \frac{\exp[-\beta U_{\text{bonded}}(i)]}{\sum_{j=1}^{n_i} \exp[-\beta U_{\text{bonded}}(j)]} = \frac{\exp[-\beta U_{\text{bonded}}(i)]}{W^{\text{int}}(n)}$$

- Repeat until  $k$  trial orientations are found; these are fed into CBMC leading to  $W(n)$
- Repeat procedure for old configuration, leading to  $W^{\text{int}}(o)$  and  $W(o)$ .
- Accept or reject according to

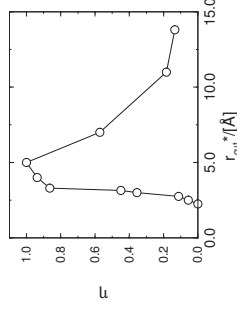
$$\text{acc}(o \rightarrow n) = \min \left( 1, \frac{W(n) \times W^{\text{int}}(n)}{W(o) \times W^{\text{int}}(o)} \right)$$

## Application 1: Adsorption of Alkanes in MFI-type zeolite (1)

Zeolites:

- microporous channel structure
- crystalline,  $\text{SiO}_2$  building blocks
- substitution of  $\text{Si}^{4+}$  by  $\text{Al}^{3+}$  and a cation ( $\text{Na}^+$  or  $\text{H}^+$ )
- typical pore size: 4 – 12 Å
- synthetic and natural; >190 framework types

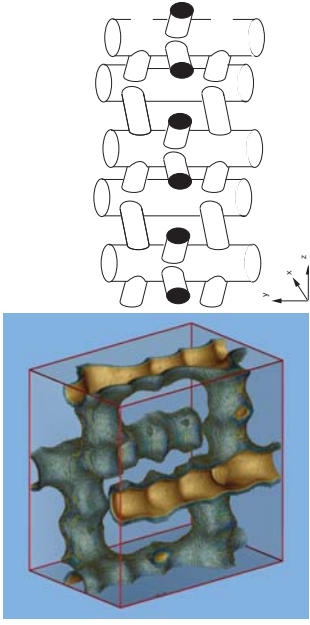
## Significant Speedup: Dual-Cutoff CBMC



- Grow chain with approximate (cheaper) potential;  $W^*$
- Correct for difference later ( $\delta u$ , difference real and approximate potential for *selected* configuration)
 
$$\text{acc}(o \rightarrow n) = \min \left( 1, \frac{W^*(n)}{W^*(o)} \times \exp[-\beta[\delta u(n) - \delta u(o)]] \right)$$



Application 1: Adsorption of Alkanes in MFI-type zeolite (2)

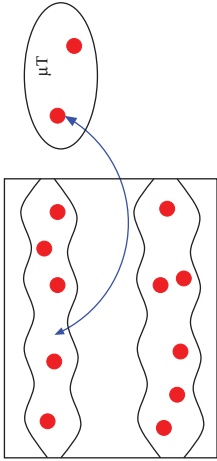


straight channels (*y* direction), zig-zag channels (*xz* plane) and intersections

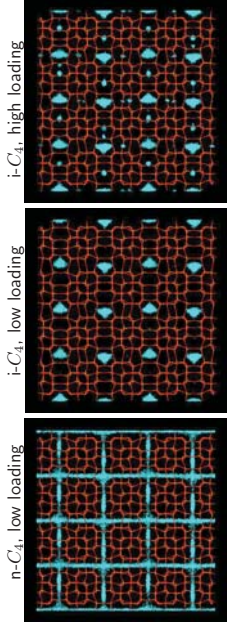
Application 1: Adsorption of Alkanes in MFI-type zeolite (3)

Grand-canonical ( $\mu VT$ ) ensemble; number of particles ( $N$ ) fluctuates

- system is coupled to particle reservoir at chemical potential  $\mu$  and temperature  $T$ , statistical weight  $\sim V^N \exp[\beta \mu N - \beta U(\mathbf{r}^N)] / (\Delta^{3N} N!)$
- trial moves to exchange particles between zeolite and reservoir (using CBMC)
- equilibrium:  $\mu_{\text{gas}} = \mu_{\text{zeolite}}$ ; measure average  $\langle N \rangle$  for given  $\mu$  and  $\beta$

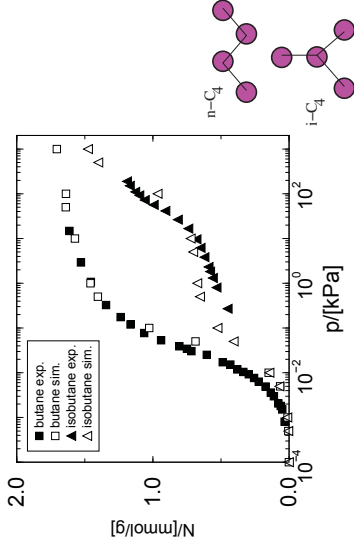


Application 1: Adsorption of Alkanes in MFI-type zeolite (5)



Vlugt et al, J. Am. Chem. Soc., 1998, 120, 5599

Application 1: Adsorption of Alkanes in MFI-type zeolite (4)

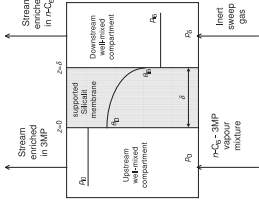


Vlugt et al, J. Am. Chem. Soc., 1998, 120, 5599

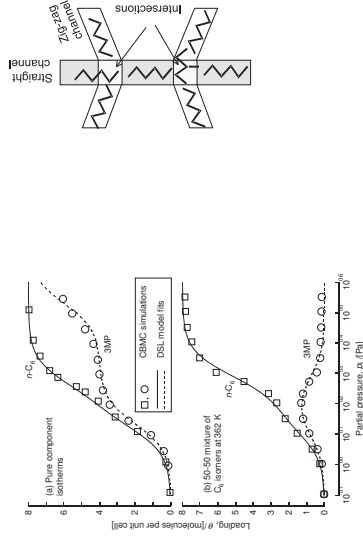
Application 1: Adsorption of Alkanes in MFI-type zeolite (7)

| Flux    | $n - C_6$ | $i - C_6$ | selectivity |
|---------|-----------|-----------|-------------|
| pure    | 179       | 136       | 1.3         |
| 50%-50% | 46        | 1.9       | 24          |

Experiments by J. Falconer, Univ. Colorado

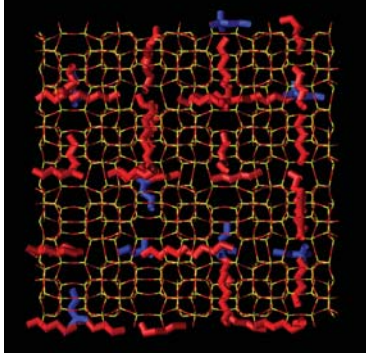


### Application 1: Adsorption of Alkanes in MFI-type zeolite (8)

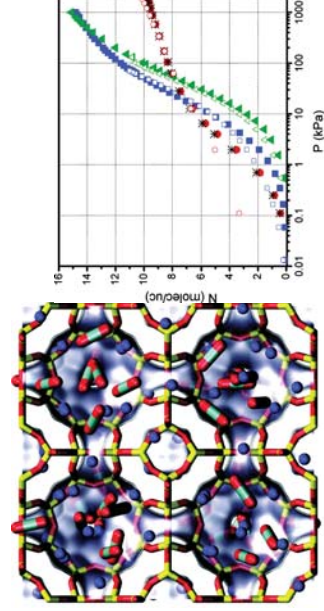


### Application 1: Adsorption of Alkanes in MFI-type zeolite (9)

blue = branched ( $i-C_6$ ) red = linear ( $n-C_6$ )

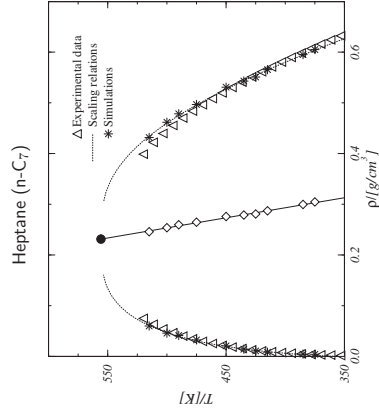


### Application 2: Adsorption of $CO_2$ in $Na^+$ containing zeolites



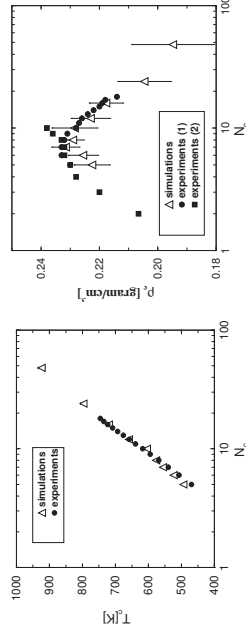
Sofia Calero et al., J. Phys. Chem. C, 2009, 113, 8814-8820

### Application 3: Gibbs Ensemble Monte Carlo (1)



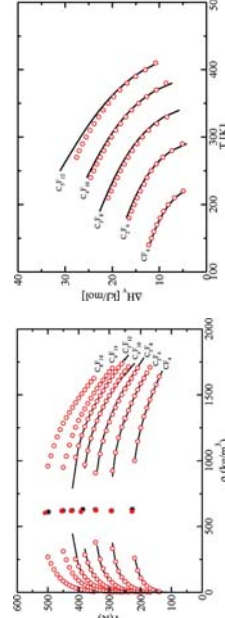
B. Smit, S. Karaborni, and J.I. Siepmann, J. Chem. Phys. 102, 2126 (1995)

### Application 3: Gibbs Ensemble Monte Carlo (2)



B. Smit, S. Karaborni, and J.I. Siepmann, J. Chem. Phys. 102, 2126 (1995)

### Application 4: GCMC Histogram Reweighting



J.J. Potoff et al., J. Phys. Chem. B, 2009, 113, pp 14725-14731

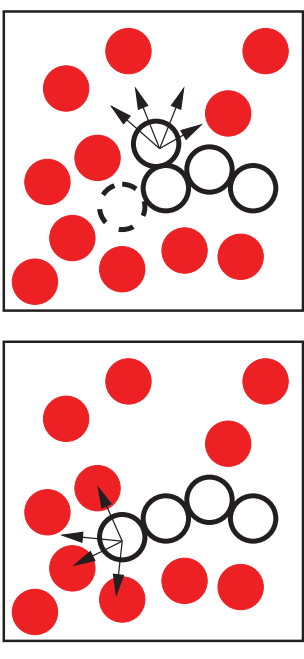
## Efficiency of CBMC

- $k$ : number of trial directions
  - $\alpha$ : probability that trial direction has a “favorable” energy
  - growth can continue as long as at least 1 trial direction is “favorable”
  - generate chain of length  $N$  successfully
- $$P_{\text{success}} = (1 - (1 - \alpha)^k)^N = \exp[-cN]$$
- increasing  $k$  means increasing CPU time.

## Dead-End Alley



## Recoil Growth: Avoiding Dead-End Alley

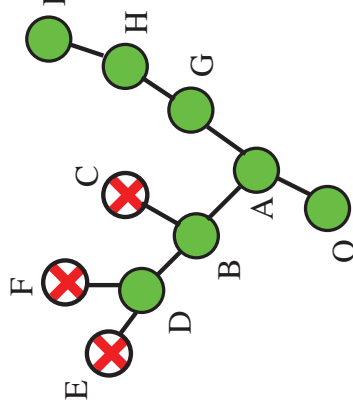


## Recoil Growth Algorithm (1)

- Place first bead at random position
- Position ( $i$ ) can be “open” or “closed” depending on the environment (energy  $u_i$ ); here we use  $p_{\text{open}} = \min(1, \exp[-\beta u_i])$  and toss a coin.
- If “open”, continue with next segment
- If “closed”, try another trial direction up to a maximum of  $k$
- If all  $k$  directions are closed, retract by one step
- Maximum retraction length:  $l_{\text{max}} - l + 1$
- $l$ : recoil length,  $l_{\text{max}}$ : maximum length obtained during the growth of the chain
- Computed weight  $W(n)$  and repeat procedure for old configuration
- Accept or reject using  $\text{acc}(o \rightarrow n) = \min(1, W(n)/W(o))$

## Recoil Growth Algorithm (2)

Example for  $k = 2$  and  $l = 3$



## Super Detailed Balance

In general,

$$N(o) \times \alpha(o \rightarrow n) \times \text{acc}(o \rightarrow n) = N(n) \times \alpha(n \rightarrow o) \times \text{acc}(n \rightarrow o)$$

Therefore,

$$\text{acc}(o \rightarrow n) = \min \left( 1, \exp[-\beta \Delta U] \times \frac{\alpha(n \rightarrow o)}{\alpha(o \rightarrow n)} \right)$$

What about  $\alpha(o \rightarrow n)$  ?

- Generate a tree  $t_n$ .
- Decide which parts of the tree are “open” or “closed” ( $O_n$ ).
- Make a random walk on the tree ( $ru_n$ )

